

10.4: Probabilistic Reasoning: Rules of Probability

- Inductive logic involves the notion of *strength* in its definition:

Inductive logic is the part of logic that is concerned with the study of methods of evaluating arguments for *strength* or *weakness*.

- And *strength* in turn was characterized in the last lecture in terms of *probability*:

A **strong argument** is one in which it is probable (but not necessary) that if the premises are true, then the conclusion is true.

- A sound foundation for inductive logic therefore requires a rigorous theoretical understanding of the notion of probability.
- And, in fact, *probability theory* has become an extremely advanced branch of mathematics.
- In these final lectures we will study the basic laws, or rules, of the *probability calculus*, which form the basis of probability theory.

Background to Probability

- There is serious philosophical disagreement about the precise nature of probability
 - Is it something “objective”, something to be *discovered* out in the world?
 - Is it just a measure of one’s own subjective feelings, a measure of the strength of one’s belief that something will occur?

- But there is widespread agreement about (a) the probabilities of certain logically distinctive propositions and (b) how the probability of a *compound* statement is determined by the probabilities of its component statements.
- The *probability calculus* consists of the basic rules concerning (a) and (b).
- The rules concerning (b) are analogous to the rules of the truth table method of Ch. 7.
 - A truth table does not tell us the truth value of simple statements like F and G.
 - But it *does* tell us how the truth value of a compound statement like $(F \vee G)$ is determined, *given* the truth values of F and G.
 - Likewise, the probability calculus does not tell us the probability of simple statements like F and G.
 - But it *does* tell us how the probability of a compound statement like $(F \vee G)$ — written $P(F \vee G)$ — is determined, *given* the probabilities of F and G — written $P(F)$ and $P(G)$.
- We will be using the language of statement logic except that the letter “P” will be reserved for the probability operator.
 - Statement letters: A, B, C, ..., O, Q, R, ..., Z (though I’ll use some alternatives below)
 - Logical operators: \sim , \cdot , \vee , \rightarrow , \leftrightarrow , and P
 - And, as before, we will use lowercase italic letters p, q, r, \dots as *metavariables* that stand for arbitrary statements.

The Rules of Probability

- Probability values are expressed as numbers from 0 to 1.
 - 0 is the lowest degree of probability, 1 the highest.

- It is customary to assign a probability of 1 to the tautologies of statement logic, i.e., those that are true in every row of a truth table.
 - This is reasonable because tautologies *must* be true; there is not the smallest probability that a tautology could be false.
- This is in fact the first rule of the probability calculus:

Rule 1: If a statement p is a tautology, then $P(p) = 1$.

- Likewise, a probability of 0 is assigned to contradictions, i.e., those that are false in every row of a truth table.

Rule 2: If a statement p is a contradiction, then $P(p) = 0$.

Examples

- By Rule 1, $P(A \vee \sim A) = P(B \rightarrow (A \rightarrow B)) = 1$.
- By Rule 2, $P(A \cdot \sim A) = P(\sim(B \vee B)) = 0$.

MUTUAL EXCLUSIVITY

- Consider the statements:
 - (a) Bernie Sanders will win the US presidency in 2024.
 - (b) RFK Jr. will win the US presidency in 2024.
- These statements both have a probability between 0 and 1.
- However they cannot both be *true*; they are *mutually exclusive*.

Two statements are **mutually exclusive** if they cannot both be true.

EXHAUSTIVENESS

- Consider the statements:
 - (a) W. V. Quine was born before 1900.
 - (b) W. V. Quine was born after 1900.
 - (c) W. V. Quine was not born before or after 1900.
- Not only are they mutually exclusive, one of them must be true; together they *exhaust* the possibilities. Hence:

Statements p, q, r, \dots are **jointly exhaustive** if at least one of them must be true.

- Now suppose p and q are mutually exclusive.
 - Let T = The die will turn up 3
 - Let S = The die will turn up 6
 - There is a 1 in 6 ($1/6$) chance that the die will land on any given side.
 - So $P(T) = P(S) = 1/6$
 - Hence, since T and S are mutually exclusive, there is a 2 in 6 chance that *either* T or S , that is:
 - $P(T \vee S) = P(T) + P(S) = 2/6 = 1/3$.
- This illustrates the *restricted disjunction rule*:

Rule 3: If p and q are mutually exclusive, then
$$P(p \vee q) = P(p) + P(q).$$

Examples

- Suppose we want to draw one card from a well-shuffled deck of 52.
- Since drawing an Ace of Clubs (**A♣**) and drawing an Ace of Diamonds (**A♦**) are mutually exclusive, we have:

$$P(\mathbf{A}\clubsuit \vee \mathbf{A}\diamondsuit) = P(\mathbf{A}\clubsuit) + P(\mathbf{A}\diamondsuit) = 1/52 + 1/52 = 2/52 = 1/26$$

- What is the probability of drawing a Queen (of any suit)?

$$P((\mathbf{Q}\clubsuit \vee \mathbf{Q}\diamondsuit) \vee (\mathbf{Q}\heartsuit \vee \mathbf{Q}\spadesuit)) = P(\mathbf{Q}\clubsuit) + P(\mathbf{Q}\diamondsuit) + P(\mathbf{Q}\heartsuit) + P(\mathbf{Q}\spadesuit) = 1/52 + 1/52 + 1/52 + 1/52 = 4/52 = 1/13$$

THE PROBABILITY OF NEGATIONS

- The restricted disjunction (RD) rule enables us to calculate the probability of a negation, $P(\sim p)$, from the probability of the statement negated, $P(p)$.
- Consider any statement p .
- p and its negation $\sim p$ are mutually exclusive.
- Hence, by the RD rule

$$P(p \vee \sim p) = P(p) + P(\sim p)$$

- But by Rule 1, the rule for tautologies, we also know that

$$P(p \vee \sim p) = 1$$

- Putting these two together, we have

$$P(p) + P(\sim p) = 1$$

- And, subtracting $P(p)$ from both sides, we have our fourth rule, the *negation rule*:

Rule 4: $P(\sim p) = 1 - P(p)$

Example 1

- Suppose we know that the probability, $P(F)$, of throwing a 4 on the next throw of a die is 1 in 6, so $P(F) = 1/6$.
- Then the negation rule enables us to calculate the probability $P(\sim F)$ that a 4 will *not* turn up on the next throw:

$$P(\sim F) = 1 - P(F) = 1 - 1/6 = 6/6 - 1/6 = 5/6.$$

Example 2

- Since there are 13 cards in each suit, the probability, $P(S)$, that we will draw a spade from a well-shuffled deck is $13/52$.
- Hence, the probability $P(\sim S)$ that we will *not* draw a spade is:

$$P(\sim S) = 1 - P(S) = 1 - 13/52 = 52/52 - 13/52 = 39/52 = 3/4.$$

THE GENERAL DISJUNCTION RULE

- Obviously, not every pair of statements is mutually exclusive.
 - In many cases p and q can both be true.
 - E.g., Let $K = \text{You draw a King}$ and $C = \text{You draw a Club}$. K and C are not mutually exclusive because of the King of Clubs.
- So we need a more *general* disjunction rule for calculating probabilities $P(p \vee q)$ when p and q are not mutually exclusive.
- Consider the probability $P(K \vee C)$ of drawing a King or a Club.
 - The sum $P(K) + P(C) = 4/52 + 13/52 = 17/52$ is too high, since we are in effect counting **K♣** twice — once as a King and once as a Club.
 - So we need *subtract* the probability of drawing **K♣**, i.e., the probability $P(K \cdot C)$ of drawing *both* a King and a club:

$$P(K \vee C) = P(K) + P(C) - P(K \cdot C) = 4/52 + 13/52 - 1/52 = 16/52 = 4/13$$

- This illustrates the *general disjunction rule*:

$$\text{Rule 5: } P(p \vee q) = P(p) + P(q) - P(p \cdot q)$$

- Note that, when p and q are mutually exclusive, $P(p \cdot q) = 0$.
- Hence, we can *derive* Rule 3 from Rule 5.
 - E.g., since it is impossible to draw *both* a Club (\clubsuit) *and* a Diamond (\diamond) on a single draw, the probability of doing so, $P(\clubsuit \cdot \diamond)$, is 0. Hence:

$$P(\clubsuit \vee \diamond) = P(\clubsuit) + P(\diamond) - P(\clubsuit \cdot \diamond) = 1/4 + 1/4 - 0 = 2/4 = 1/2.$$

Example

- What is the probability $P(R \vee E)$ of drawing a red card (R) or an 8 (E)?
 - R = You draw either a Heart or a Diamond, ($\heartsuit \vee \diamond$).
 - So $P(R) = P(\heartsuit \vee \diamond) = P(\heartsuit) + P(\diamond)$ (since \heartsuit and \diamond are mutually exclusive) = $13/52 + 13/52 = 26/52 (= 1/2)$.
 - E = You draw either $8\clubsuit, 8\diamond, 8\heartsuit$, or $8\spadesuit$
 - So $P(E) = P(8\clubsuit \vee 8\diamond \vee 8\heartsuit \vee 8\spadesuit) = P(8\clubsuit) + P(8\diamond) + P(8\heartsuit) + P(8\spadesuit) = 4/52 = 1/13$.
 - Since there are two red eights, $8\diamond$ and $8\heartsuit$, $P(R \cdot E) = 2/52$.
- $P(R \vee E) = P(R) + P(E) - P(R \cdot E) = 26/52 + 4/52 - 2/52 = 28/52 = 7/13$.

CONDITIONAL PROBABILITY

- Because $p \rightarrow q$ is logically equivalent to $\sim p \vee q$ (recall the MI rule), it follows that $P(p \rightarrow q) = P(\sim p \vee q)$.
- *But*, as I've noted before, the meaning we've assigned to \rightarrow (via its truth table) does not adequately capture the meaning of "if ... then" in every context — notably, those involving judgments of probability.

- Consequently, a rule of probability has been designed to capture the meaning of conditionals in such contexts.
- Specifically, this rule is designed to enable us to calculate the probability that q is true *conditional on* p 's being true.
 - We will write “The probability of q conditional on p ” as $P(q/p)$.
 - This notation can also read as:
 - The probability of q on the condition that p .
 - The probability of q on p
 - The probability of q given p .”
 - In statements of the form $P(q/p)$, p is the *antecedent* and q the *consequent*.
- The *conditional rule* is as follows:

Rule 6:

$$P(q/p) = \frac{P(p \cdot q)}{P(p)}$$

- Why was it decided that $P(q/p)$ is the $P(p \cdot q)$ divided by $P(p)$?

Example 1

- Suppose we are about to draw one card from a well-shuffled deck.
- What's $P(\clubsuit/\spadesuit)$, i.e., the probability of our drawing a club *given that* we will draw \spadesuit ?
 - Intuitively, it is certain, i.e., it should turn out that $P(\clubsuit/\spadesuit) = 1$.
- $P(\clubsuit/\spadesuit) = P(\clubsuit \cdot \spadesuit)/P(\spadesuit) = P(\spadesuit)/P(\spadesuit) = 1$.

Example 2

- What's $P(\spadesuit/\heartsuit)$, i.e., the probability of our drawing a Spade *given that* we will draw a Heart?
 - Intuitively, it is nil, i.e., it should turn out that $P(\spadesuit/\heartsuit) = 0$. For, *given* we will draw a Heart, we can't possibly draw another suit.
- $P(\spadesuit/\heartsuit) = P(\heartsuit \cdot \spadesuit)/P(\heartsuit) = 0/P(\heartsuit) = 0/1/4 = 0$.

Example 3

- What's $P(\heartsuit/K)$, i.e., the probability of our drawing a King of Hearts *given that* we will draw a King (of any suit).
 - Intuitively, it should be $1/4$. For, *given* that we will draw a King, there is a 1 in 4 chance that it will be the King of Heart instead of one of the other three.
- $P(\heartsuit/K) = P(K \cdot \heartsuit)/P(K) = P(\heartsuit)/P(K) = \frac{1/52}{4/52} = 1/52 \times 52/4 = 1/4$.

Example 4

- What's $P(\clubsuit/\clubsuit \vee \spadesuit)$, i.e., the probability of our drawing a Club *given that* we will draw black card, i.e., either a Club or a Spade?
 - Intuitively, it should be $1/2$. For, *given* that we will draw a black card, it must be either a Club or a Spade. Since the number of Clubs = the number of Spades, there is a 1 in 2 chance our card will be a Club.
- $P(\clubsuit/\clubsuit \vee \spadesuit) = P((\clubsuit \vee \spadesuit) \cdot \clubsuit)/P(\clubsuit \vee \spadesuit) = P(\clubsuit)/P(\clubsuit \vee \spadesuit) = \frac{13/52}{26/52} = \frac{1/4}{1/2} = \frac{1}{4} \times \frac{2}{1} = 1/2$.

CONJUNCTION

- The conditional rule is important, not only for what it tells us about conditional probability but also because from it we can immediately deduce the *general conjunction rule*:

Rule 7: $P(p \cdot q) = P(p) \times P(q/p)$
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- To prove this, note that by the conditional rule (Rule 6) we have:

$$P(q/p) = \frac{P(p \cdot q)}{P(p)}$$

- Next, we multiply both sides of the equation by $P(p)$:

$$P(p) \times P(q/p) = P(p) \times \frac{P(p \cdot q)}{P(p)}$$

- Since $a \times \frac{b}{a} = \frac{a}{1} \times \frac{b}{a} = \frac{1}{1} \times \frac{b}{1} = b$ we have:

$$P(p) \times P(q/p) = P(p \cdot q)$$

- And that is exactly the conjunction rule (with the two sides switched).

Example 1

- Consider the situation where you draw a card and then, without replacing the first card, draw a second card.
- Let $\mathbf{A\spadesuit}_1$ be drawing $\mathbf{A\spadesuit}$ on the first draw and $\mathbf{A\spadesuit}_2$ be drawing $\mathbf{A\spadesuit}$ on the second draw. What is $P(\mathbf{A\spadesuit}_1 \cdot \mathbf{A\spadesuit}_2)$
- $P(\mathbf{A\spadesuit}_1 \cdot \mathbf{A\spadesuit}_2) = P(\mathbf{A\spadesuit}_1) \times P(\mathbf{A\spadesuit}_2/\mathbf{A\spadesuit}_1) = 1/52 \times 0 = 0$

Example 2

- What is the probability $P(\mathbf{Red}_1 \cdot \mathbf{Red}_2)$ of choosing a Red card (i.e., a Heart or a Diamond) and then, without putting it back, choosing another?
- $P(\mathbf{Red}_1 \cdot \mathbf{Red}_2) = P((\heartsuit \vee \diamondsuit)_1 \cdot (\heartsuit \vee \diamondsuit)_2) = P((\heartsuit \vee \diamondsuit)_1) \times P((\heartsuit \vee \diamondsuit)_2 / (\heartsuit \vee \diamondsuit)_1) = 1/2 \times 25/51 = 25/102$.

Example 3

- What is the probability $P(\mathbf{A}_1 \bullet \mathbf{A}_2)$ of drawing an ace on the first draw *and* (without replacing the first card drawn) another ace on the second draw?
- $P(\mathbf{A}_1 \bullet \mathbf{A}_2) = P(\mathbf{A}_1) \times P(\mathbf{A}_2/\mathbf{A}_1) = 4/52 \times 3/51 = 1/13 \times 1/17 = 1/221$.

INDEPENDENCE

- Our final rule requires us to introduce the important notion of *independence*.

Two statements p and q are **independent** if neither affects the probability of the other, that is, if $P(q/p) = P(q)$ and $P(p/q) = P(p)$.

Example

- *Hillary Clinton will be the next US President (H)* is independent of *The first card I choose (from a full deck) will be an Ace (A)*.
 - So $P(A/H) = P(A)$
- *The second card I choose will be a Queen (Q)* is *not* independent of *The first card I choose will be Jack (J)*.
 - In this case, $P(Q/J) = 4/51$.
- When we're dealing with independent propositions, we can derive a simpler rule for conjunctions, the *restricted conjunction rule*:

Rule 8: $P(p \bullet q) = P(p) \times P(q)$

- By Rule 7, $P(p \bullet q) = P(p) \times P(q/p)$.
- But since p and q are independent, $P(q/p) = P(q)$.

Example

- Consider the probability of selecting an ace twice by drawing from a well-shuffled deck, replacing the card, reshuffling, and drawing a second time.

$$P(A_1 \cdot A_2) = P(A_1) \times P(A_2) = 1/13 \times 1/13 = 1/169$$

An Important Observation

- The restricted conjunction rule highlights an important fact about probability.
- Suppose we have a conjunction of independent statements, each of which has a probability of less than 1 but greater than 1/2.
- For example, suppose $P(A) = P(B) = P(C) = 7/10$.
- What is the probability of the whole conjunction? Because A, B, and C are independent we have:

$$P(A \cdot B \cdot C) = P(A) \times P(B) \times P(C) = (7/10)^3 = 343/1000$$

- Although each conjunct is more probable than not, the entire conjunction has a probability of less than 1/2.
- *Bottom line:* A conjunction of likely truths can itself be unlikely.

Bayes' Theorem

- We will now focus on one important implication of our system: *Bayes' theorem*.
 - Named after the English theologian and mathematician Thomas Bayes (1702–1761).
- Bayes' theorem gives us an important insight into the relationship between the evidence for a hypothesis and the hypothesis itself, hence, it promises a deeper understanding of the scientific method.
- The letter *h* will stand for a given hypothesis.

- The letter e will stand for a statement that summarizes the observational evidence for that hypothesis.
 - Normally, e is a statement expressing the *latest* observational evidence for h
 - So Bayes' Theorem yields particular insight into the effect of a *new* piece of evidence for a hypothesis for which some body of evidence already exists.

The Derivation of Bayes' Theorem

- Bayes' Theorem is actually a surprisingly simple theorem of the probability calculus.
- We start with an instance of the conditional rule (Rule 6), for a given hypothesis h and piece of evidence e :

$$P(h/e) = \frac{P(e \cdot h)}{P(e)}$$

- A simple truth table (or proof) shows that e is logically equivalent to $(e \cdot h) \vee (e \cdot \sim h)$.
- Hence, we can replace e with $(e \cdot h) \vee (e \cdot \sim h)$ wherever we wish. Doing so in the denominator yields:

$$P(h/e) = \frac{P(e \cdot h)}{P((e \cdot h) \vee (e \cdot \sim h))}$$

- By the restricted disjunction rule (Rule 3),

$$P((e \cdot h) \vee (e \cdot \sim h)) = P(e \cdot h) + P(e \cdot \sim h)$$

- Hence:

$$P(h/e) = \frac{P(e \cdot h)}{P(e \cdot h) + P(e \cdot \sim h)}$$

- By the statement logic rule of commutation for • we have:

$$P(h/e) = \frac{P(h \cdot e)}{P(h \cdot e) + P(\sim h \cdot e)}$$

- By applying the general conjunction rule (Rule 7) three times, we arrive at **Bayes' Theorem**:

$$P(h/e) = \frac{P(h) \times P(e/h)}{[P(h) \times P(e/h)] + [P(\sim h) \times P(e/\sim h)]}$$

Implications and Applications of Bayes' Theorem

- Bayes' theorem tells us the degree to which a given hypothesis is supported by the evidence, provided that we have three pieces of information: $P(h)$, $P(e/h)$, and $P(e/\sim h)$.
 - Recall we can calculate $P(\sim h)$ from $P(h)$.
- $P(h)$ stands for the *prior probability* of the hypothesis h .

The **prior probability** of a hypothesis h is the likelihood of the hypothesis *independent of any new evidence* e .

- $P(e/h)$ is the likelihood that the evidence (or phenomenon in question) would be present, assuming the hypothesis is *true*.
- $P(e/\sim h)$ is the likelihood that the evidence (or phenomenon in question) would be present, assuming the hypothesis is *false*.

Example 1

- Suppose a doctor has diagnosed a patient as having *either* some minor stomach troubles *or* stomach cancer.
 - Let us assume as well that the doctor knows that the patient does not have *both* minor stomach troubles *and* stomach cancer.
- The doctor also knows that, given the symptoms, 30% of patients have stomach cancer; the rest have minor stomach troubles.
- The doctor initially suspects that the patient has only minor stomach troubles.
- But the doctor then conducts a test on the patient.
- The test has positive result = 90% chance of stomach cancer.
 - Let H = *the patient has stomach cancer*
 - Let E = *the test is positive*
- What is the probability of H given E , i.e., what is $P(H/E)$?
 - NOTE: You might think the obvious answer is 90% but recall that the doctor has a prior hypothesis that the patient only has a 30% chance of cancer.
- $P(H)$ = the *prior* probability of H , before E = 30% = .3 = 3/10.
- $P(\sim H)$ = 70% = .7 = 7/10 (by the negation rule, Rule 4).
- $P(E/H)$ = 90% = .9 = 9/10.
- $P(E/\sim H)$ = 10% = .1 = 1/10.
- Plugging these values directly into Bayes' Theorem, we have:

$$P(H/E) = \frac{3/10 \times 9/10}{[3/10 \times 9/10] + [7/10 \times 1/10]} = \frac{27/100}{27/100 + 7/100} = \frac{27}{34}$$

- So, the probability of the hypothesis H given the evidence E is 27/34, or approximately .79.

- I will avoid the derivation, but we note that we get an conditional analog of the negation rule (Rule 4):

$$P(\sim h/e) = 1 - P(h/e)$$

- Bayes' Theorem is still applicable when there are more than two hypothesis competing for our credence.
- If $h_1, h_2,$ and h_3 are three *mutually exclusive, jointly exhaustive* hypotheses, then $\sim h_1$ is equivalent to $h_2 \vee h_3$.
- Hence, substituting into Bayes' Theorem, we have

$$P(h_1/e) = \frac{P(h_1) \times P(e/h_1)}{[P(h_1) \times P(e/h_1)] + [P(h_2 \vee h_3) \times P(e/(h_2 \vee h_3))]}$$

- And this, in turn, reduces to

$$P(h_1/e) = \frac{P(h_1) \times P(e/h_1)}{[P(h_1) \times P(e/h_1)] + [P(h_2) \times P(e/h_2)] + [P(h_3) \times P(e/h_3)]}$$

- In other words, we can accommodate as many hypotheses as we like (provided they are mutually exclusive and jointly exhaustive), simply by adding relevant clauses to the denominator.